

Lecture 16

16 December 2024 10:11

- 6th January: Bayesian inference (Dr. Samundrajit Thapa)
- 13th January (?)
- On the NY eve: Assignment No. 3
- 20th January: lecture
- Assignment No. 4

Written exam ~ 10th February

II.4. Non-stationary processes with stationary increments

Reminder: Nonstationary processes with independent increments
Two charact. functions fully determine process

$$\hat{p}(k, t) = \langle e^{ikx(t)} \rangle$$

$$\hat{p}(k, t_1, t_2) = \langle e^{ik(x(t_2) - x(t_1))} \rangle$$

$\Downarrow$

1-point PDF and 2-point PDF

$$\hat{p}_2(k_1, t_1; k_2, t_2) = \langle e^{ik_1 x(t_1) + ik_2 x(t_2)} \rangle =$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} p_2(x_1, t_1; x_2, t_2) e^{ik_1 x_1 + ik_2 x_2} dx_1 dx_2$$

$$\hat{p}(k, t_1, t_2) = \hat{p}_2(k, t_1; -k, t_2)$$

Remark. Non-stat. processes with ind. incr. $\in$ class of Markov processes (later)

Gaussian processes with independent increments

$$\hat{p}(k, t) = \exp\left[ik\mu(t) - \frac{\sigma^2(t)k^2}{2}\right] \div p(x, t) = \frac{e^{-(x-\mu(t))^2/2\sigma^2(t)}}{\sqrt{2\pi\sigma^2(t)}}$$

$$\langle x(t) \rangle = \mu(t), \text{Var}[x(t)] = \langle (x(t) - \langle x(t) \rangle)^2 \rangle = \sigma^2(t)$$

$$\hat{p}(k, t_1, t_2) = \exp\left[ik\mu_\Delta(t_1, t_2) - \frac{\sigma_\Delta^2(t_1, t_2)k^2}{2}\right] \leftrightarrow$$

$$\leftrightarrow p(x_1, t_1; x_2, t_2) = p(x_2 - x_1, t_1, t_2) = \{\text{see previous lecture}\}$$

$$= \frac{1}{\sqrt{2\pi(\sigma^2(t_2) - \sigma^2(t_1))}} \exp\left[-\frac{(x_2 - x_1 - (\mu_2(t_2) - \mu_1(t_1)))^2}{2(\sigma^2(t_2) - \sigma^2(t_1))}\right]$$

$$t_2 > t_1$$

Particular cases

1. Wiener process (ordinary Brownian motion) (oBm)

$$\mu(t) = 0, \quad \sigma^2(t) = ct, \quad c > 0$$

2. Scaled Brownian motion (sBm)

$$\mu(t) = 0, \quad \sigma^2(t) = ct^\alpha, \quad c > 0, \alpha > 0, \alpha \neq 1$$

(4.1) Definition nonstat. process with stat. incr.

- $x(t)$ nonstationary, but

- $x(t+\tau) - x(t)$ stationary in t
(statistical properties of $x(t_2) - x(t_1)$
depend on $t_2 - t_1$), but:

can be non-independent $\Rightarrow \hat{p}(k, t)$ and $\hat{p}(k, t_1, t_2)$
are not enough (in general) to define $x(t)$.

Remark. If $x(t)$ stationary $\Rightarrow$ increments of $x(t)$
are stationary as well

(4.2) Mean value

Remind: for stationary process $\langle x(t) \rangle = \text{const}$

$\langle x(t+\tau) - x(t) \rangle \neq f(t) \Rightarrow$ e.g., $\langle x(t) \rangle = at + b$,

Example 1. Wiener process (oBm): $a, b = \text{const}$

increments independent and stationary

$$\sigma^2(t+\tau) - \sigma^2(t) = c\tau$$

Example 2. Gaussian with $\sigma^2(t) = ct$,

linear drift $\mu(t) = \nu t$, $\nu = \text{const} \Rightarrow$

$\Rightarrow$ increments independent and stationary

Example 3. Scaled Brownian motion (sBm) $\Rightarrow$

$\Rightarrow$ increments independent but nonstationary

Example 4. $x(t) = \alpha t + \xi(t)$

α r.v. $\langle \alpha \rangle = a$, $\xi(t)$ stationary, $\langle \xi(t) \rangle = b$

$\Rightarrow$ increments are stationary

4.3) Structure function $D(\tau)$

$$x(t) = \langle x(t) \rangle + \xi(t), \quad \langle \xi(t) \rangle = 0$$

↑ fluctuations of $x(t)$

$$D(\tau) = \langle (\xi(t+\tau) - \xi(t))^2 \rangle$$

- Structure function $\equiv$ mean squared increment, MSI
- Analogue of the corr function
- Can be also constructed for stationary process

$$\begin{aligned} D(\tau) &= \langle \xi^2(t+\tau) \rangle - 2\langle \xi(t)\xi(t+\tau) \rangle + \langle \xi^2(t) \rangle = \\ &= \{ B_{\xi}(\tau) = \langle \xi(t)\xi(t+\tau) \rangle \} = \\ &= 2 [B_{\xi}(0) - B_{\xi}(\tau)] \end{aligned}$$

Remark. $B(\tau) \leq B(0)$ because $D(\tau) \geq 0$

- Can be evaluated as a time average along a single trajectory

$$\langle (\xi(t+\tau) - \xi(t))^2 \rangle = \frac{1}{T-\tau} \int_0^{T-\tau} (x(t+\tau) - x(t))^2 dt$$

$T \rightarrow \infty$

Example. Random line $x(t) = \alpha t + c$
 α r.v., $c = \text{const}$

$$D(\tau) = A\tau^2, \quad A = \langle \alpha^2 \rangle$$

- upper bound for structure function of the process with stationary increments

$$D(\tau) \leq A\tau^2, \quad \tau \uparrow$$

(proved with Minkowski's inequality)

- class of scale-free structure functions (anomalousity) (Kolmogorov in 1941)

$$D(\tau) = A\tau^{2H}, \quad 0 < H < 1 \quad \text{hurst exponent (physics lit.)}$$

~~Special symmetry in physical system:~~

scale character of independent variables (time...)

Special symmetry in physical system:
scale changes of independent variable (time)
are compensated with similar transformations
of another dynamical variable

$$t \rightarrow \lambda t, \quad x \rightarrow \lambda^{-H} x \Rightarrow D(\lambda \tau) = \lambda^{2H} D(\tau)$$

- Definition: self-similar random process or random fractal

$$X(\lambda t) \stackrel{d}{=} \lambda^H X(t) \quad \stackrel{d}{=}: \text{equality in probability}$$

Check. lhs and rhs have the same structure function.

II.5. Lévy processes: independent and stationary increments
(homogeneous processes with independent increments)

$$(5.1) \quad X(t+\tau) - X(t) \stackrel{d}{=} X(\tau)$$

$$X(0) = 0$$

Important: PDF characterized by its CF
(Lévy-Khinchine formula)

$$\hat{p}(k, t) = e^{\underbrace{t\Psi(k)}}_{\substack{\uparrow \\ \text{L-Kh establish general form of } \Psi(k)}}$$

Example: oBm with linear drift

$$\Psi(k) = ik\mu - \frac{1}{2}\sigma^2 k^2$$

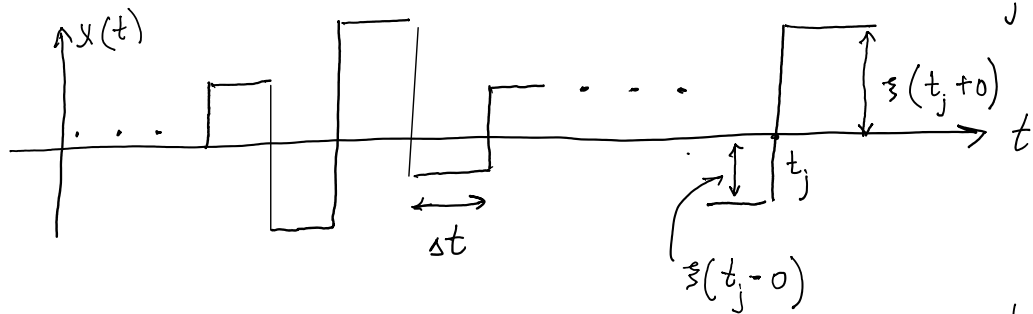
(5.2.) continuous-time analogue of sums of i.i.d. r.v

$$X(t_n) = X(t_1) + \underbrace{\sum_{j=2}^n [X(t_j) - X(t_{j-1})]}_{\substack{\text{random, independent,} \\ \text{statistically identical over} \\ \text{time intervals of the same length} \\ \text{(because of their stationarity)}}}$$

random, independent,
statistically identical over
time intervals of the same length
(because of their stationarity)

Example. Random process with iid random jumps in fixed time intervals

$$\text{jumps } \xi_j = \xi(t_j + 0) - \xi(t_j - 0), \quad x(t) = \sum_{t_j < t} \xi_j$$



With this toy example we built two important continuous-time processes

(5.3) Wiener process

a) discrete time representation

$$\text{setup } t_1 = 0, \quad t_n = t, \quad \Delta t = \frac{t}{n}, \quad x(t) = \sum_{j=1}^n \xi_j$$

$$\text{jumps } \xi_j \text{ Gaussian } \hat{p}_\xi(k) = e^{-\frac{1}{2}\sigma^2 k^2}$$

$$\updownarrow$$

$$p_\xi(x) = \frac{e^{-x^2/2\sigma^2}}{\sqrt{2\pi\sigma^2}}, \quad \langle \xi^2 \rangle = \sigma^2, \quad \langle \xi \rangle = 0$$

$$\begin{aligned} \bullet \text{ PDF } x(t): \quad p(x, t) &\div \hat{p}(k, t) = [\hat{p}_\xi(k)]^n = e^{-\frac{\sigma^2}{2} k^2 n} = \\ &= \left\{ n = \frac{t}{\Delta t} \right\} = \exp\left[-\frac{1}{2} D k^2 t\right], \quad D = \frac{\sigma^2}{\Delta t}, \quad \Psi(k) = \frac{D k^2}{2} \end{aligned}$$

$$p(x, t) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx} \hat{p}(k, t) dk = \frac{\exp\left(-\frac{x^2}{2Dt}\right)}{\sqrt{2\pi Dt}}$$

$$\text{Check: } \int_{-\infty}^{\infty} p(x, t) dx = 1, \quad \langle x(t) \rangle = 0, \quad \langle x^2(t) \rangle = Dt$$

Conclusion 1. Jumps are Gaussian iid $\Rightarrow$

$$\Rightarrow p(x, t) \text{ Gaussian with } \text{Var}[x(t)] = Dt$$

Q. how CLT is applied if $\{\xi_j\}$ are not Gaussian,
 BUT $\langle \xi \rangle = 0$, $\langle \xi^2 \rangle = \sigma^2 < \infty \Rightarrow$
 $\Rightarrow p_\xi(x)$ belongs to the domain of attraction of
 the Gaussian law

Reminder. CLT Properly normalized sum

$$Z = \frac{1}{\sqrt{n}} (\xi_1 + \dots + \xi_n)$$

$$p_Z(z) \xrightarrow{n \rightarrow \infty} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{z^2}{2\sigma^2}\right) \div e^{-\frac{1}{2}\sigma^2 k^2}$$

$$\langle Z^2 \rangle = \sigma^2$$

PDF of $X = \sqrt{n} Z$

Reminder. $X = aZ \Rightarrow \hat{p}_X(k) = \hat{p}_Z(ak)$

$$p(x,t) \div \hat{p}(k,t) \xrightarrow{n \rightarrow \infty} \hat{p}_Z(\sqrt{n}k) = \exp\left(-\frac{\sigma^2}{2} k^2 n\right) = \left\{ n = \frac{t}{\Delta t} \right\}$$

$$= \exp(-D k^2 t), \quad D = \frac{\sigma^2}{2\Delta t}, \quad t \rightarrow \infty$$

$$p(x,t) = \frac{1}{\sqrt{4\pi Dt}} \exp\left(-\frac{x^2}{4Dt}\right) \quad \text{norm, } \langle x(t) \rangle = 0,$$

$$\langle x^2(t) \rangle = 2Dt$$

Conclusion 2. Jumps are now Gaussian iid with
 finite variance $\Rightarrow$ after many steps
 $p(x,t) = \text{Gauss with } \text{Var}[x(t)] = 2Dt$

Terminology. D diffusion coeff.

$$\langle (x(t) - x(0))^2 \rangle \quad \text{mean squared displacement}$$

" "
0

$\Rightarrow$ Wiener process (oBm) appears as the limit one
 for sum of iid random jumps with finite variance
 due to CLT $\Rightarrow$ important and ubiquitous
 (math rigorous formulation: Donsker Theorem)

~~due to CLT $\Rightarrow$ important and non-trivial~~
 (math rigorous formulation: Donsker Theorem)

$\Rightarrow$ Generic example of random walk: the process can be viewed as a succession of random jumps

Remark. Trajectories of Wiener process are continuous, but nondifferentiable

b) self-similarity of Wiener process $W(t)$

$$W(\lambda t) \stackrel{d}{=} \lambda^H W(t) \quad (*) \text{ Hurst}$$

$$\text{Use } \langle e^{ikW(t)} \rangle = e^{-Dk^2 t}$$

$$\left. \begin{array}{l} \text{lhs of } (*) \langle e^{ikW(\lambda t)} \rangle = e^{-Dk^2 \lambda t} \\ \text{rhs of } (*) \langle e^{ik\lambda^H W(t)} \rangle = e^{-D\lambda^{2H} k^2 t} \end{array} \right\} \Rightarrow \boxed{H = \frac{1}{2}}$$

c) Mean squared displacement (MSD) and autocorrelation function (ACF)

$$\bullet \langle W^2(t) \rangle = - \frac{\partial^2 \hat{p}(k, t)}{\partial k^2} \Big|_{k=0} = 2Dt$$

$$\bullet \langle W(t_1)W(t_2) \rangle_{t_1 < t_2} = \langle W(t_1)[W(t_1) + W(t_2) - W(t_1)] \rangle =$$

$$= \langle W^2(t_1) \rangle + \langle W(t_1)[W(t_2) - W(t_1)] \rangle = 2Dt_1$$

"
 $2Dt_1$

" independence of increments

$$\text{By analogy } \langle W(t_1)W(t_2) \rangle_{t_1 > t_2} = \langle W^2(t_2) \rangle = 2Dt_2$$

$$\boxed{\langle W(t_1)W(t_2) \rangle = 2D \min\{t_1, t_2\}}$$

$$\bullet \text{ structure function } S(\tau) = \langle (W(t+\tau) - W(t))^2 \rangle = ?$$

d) oBm as integral of white noise

Digression. More general case: integral of a stationary process

$$\frac{dx}{dt} = \xi(t), \quad x(t) = \int_0^t \xi(t') dt' \quad \begin{array}{l} \text{stationary} \\ \text{process} \end{array} \quad \begin{array}{l} \text{stochastic} \\ \text{integral} \end{array}$$

$$\langle \xi \rangle = 0, \quad \langle \xi(t) \xi(t+\tau) \rangle = B(\tau)$$

Property 1. $x(t)$ has stationary increments.

$$\frac{dx}{dt} = \lim_{\Delta t \rightarrow 0} \frac{x(t+\Delta t) - x(t)}{\Delta t}$$

Property 2 MSD for integral of stationary process

$$\begin{aligned} \langle x^2(t) \rangle &= \left\langle \int_0^t \xi(t_1) dt_1 \int_0^t \xi(t_2) dt_2 \right\rangle = \\ &= \int_0^t dt_1 \int_0^t dt_2 \langle \xi(t_1) \xi(t_2) \rangle = \int_0^t dt_1 \int_0^t dt_2 B(t_2 - t_1) = \end{aligned}$$

$$= \{ \text{remind lecture on ergodicity} \} =$$

$$= \int_0^t dt_1 \int_0^{t_1} dt_2 B(t_2 - t_1) + \int_0^t dt_1 \int_{t_1}^t dt_2 B(t_2 - t_1) =$$

change order of integration and use $B(t_2 - t_1) = B(t_1 - t_2)$

$$= 2 \int_0^t dt_1 \int_{t_1}^t dt_2 B(t_2 - t_1) = \{ \tau = t_2 - t_1 \} = 2 \int_0^t dt_1 \int_0^{t-t_1} d\tau B(\tau) =$$

change order of integr.

$$= 2 \int_0^t (t - \tau) B(\tau) d\tau, \quad \begin{array}{l} \text{see details} \\ \text{in ergodicity lecture} \end{array}$$

$$\frac{dx}{dt} = \xi(t) \text{ stationary} \Rightarrow \langle x^2(t) \rangle = 2 \int_0^t (t - \tau) B(\tau) d\tau$$

- Back to white noise $\xi(t)$

$$D = \int_0^{\infty} B(\tau) d\tau \neq \int_0^{\infty} \delta(\tau) d\tau = 1/2$$

$$2) \langle x^2(t) \rangle = 2 \int_0^{\infty} (t-\tau) B(\tau) d\tau = 4D \int_0^{\infty} (t-\tau) \delta(\tau) d\tau = 2Dt = \text{Var}[x(t)]$$

$$\begin{aligned} \text{From 1)-3)} \Rightarrow p(x,t) &= \frac{1}{\sqrt{2\pi \text{Var}[x(t)]}} \exp\left[-\frac{x^2}{2 \text{Var}[x(t)]}\right] = \\ &= \frac{1}{\sqrt{4\pi Dt}} \exp\left(-\frac{x^2}{4Dt}\right) \quad \text{Wiener (oBm)} \end{aligned}$$

$$\begin{aligned}
 \langle x(t_1) x(t_2) \rangle &= \left\langle \int_0^{t_1} dt'_1 \xi(t'_1) \int_0^{t_2} dt'_2 \xi(t'_2) \right\rangle = \\
 &= \int_0^{t_1} dt'_1 \int_0^{t_2} dt'_2 \langle \xi(t'_1) \xi(t'_2) \rangle = \left\{ \text{Let } t_2 > t_1 \right\} = \\
 &= \int_0^{t_1} dt'_1 \int_0^{t_1} dt'_2 \langle \xi(t'_1) \xi(t'_2) \rangle + \int_0^{t_1} dt'_1 \int_{t_1}^{t_2} dt'_2 \langle \xi(t'_1) \xi(t'_2) \rangle = \\
 &= 2D \underbrace{\int_0^{t_1} dt'_1 \int_0^{t_1} dt'_2 \delta(t'_2 - t'_1)}_{=1} + 2D \underbrace{\int_0^{t_1} dt'_1 \int_{t_1}^{t_2} dt'_2 \delta(t'_2 - t'_1)}_{=0} =
 \end{aligned}$$

$$\left. \begin{array}{l} 2Dt_1, \quad t_1 < t_2 \\ \text{Similarly, } \langle x(t_1)x(t_2) \rangle = 2Dt_2, \quad t_2 < t_1 \end{array} \right\} \Rightarrow$$

$$\Rightarrow \langle x(t_1) x(t_2) \rangle = 2D \min \{t_1, t_2\} \quad \text{see above}$$

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$$\text{Let } t_1 < t_2 \quad \langle x(t_1) [x(t_2) - x(t_1)] \rangle =$$

$$= \langle x(t_1) x(t_2) \rangle - \langle x^2(t_1) \rangle = 0$$

The same for $t_1 < t_2 < t_3$

$$\langle [x(t_2) - x(t_1)] [x(t_3) - x(t_2)] \rangle = \text{open the brackets} = 0$$

Remark 3 In physics lit. re often

$$\text{instead of } \frac{dx}{dt} = \xi(t), \quad \langle \xi(t) \xi(t') \rangle = 2D \delta(t - t')$$

$$\frac{dx}{dt} = \sqrt{2D} \eta(t), \quad \langle \eta(t) \eta(t') \rangle = \delta(t - t')$$

e) discrete time representation of white Gaussian noise

$$\frac{dW}{dt} = \xi(t), \quad t_{j+1} - t_j = \Delta t$$

$$W(t_{j+1}) - W(t_j) = \int_{t_j}^{t_{j+1}} \xi(t') dt' = ?$$

$$W(\Delta t) - W(0) = W(\Delta t) = (\Delta t)^{1/2} W(1) \quad (*,*)$$

self-sim. $W(\lambda t) = \lambda^H W(t)$, $H = 1/2$

$$p_W(x, t) = \frac{e^{-x^2/4Dt}}{\sqrt{4\pi Dt}} \div \hat{p}_W(k, t) = \exp(-Dk^2 t)$$

$$W(1) \equiv \xi, \quad p_\xi(x) = \frac{e^{-x^2/4D}}{\sqrt{4\pi D}} \div \hat{p}_\xi(k) = e^{-Dk^2}$$

$$\text{Let } \xi = \sqrt{2D} \eta \Rightarrow \{ \hat{p}_\xi(k) = \hat{p}_\eta(\sqrt{2D} k) \} \Rightarrow$$

$$\Rightarrow \hat{p}_\eta(k) = e^{-k^2/2} \div p_\eta(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$$

η : standard normal variable $\eta \sim \mathcal{N}(0, 1)$, $\langle \eta \rangle = 0$
 Back to (**) $\langle \eta^2 \rangle = 1$

$$W(\Delta t) = (\Delta t)^{1/2} \xi = \sqrt{2D\Delta t} \eta$$

$$\boxed{\begin{aligned} W(t_{j+1}) - W(t_j) &= (2D\Delta t)^{1/2} \eta_j \\ \eta_j &\sim \mathcal{N}(0, 1) \end{aligned}} \Leftrightarrow$$



$$\boxed{\begin{aligned} \frac{dW}{dt} &= \xi(t) = \sqrt{2D} \eta(t) \\ \langle \xi(t) \xi(t') \rangle &= 2D \delta(t - t') \\ \langle \eta(t) \eta(t') \rangle &= \delta(t - t') \end{aligned}}$$

$$\Leftrightarrow \left. \begin{aligned} dW &= \xi(t) dt \\ \Downarrow \\ W(t_{j+1}) - W(t_j) &= \xi_j \Delta t \end{aligned} \right\} ?$$

Remark 1. Math. definition of white Gaussian noise

$\xi(t)$ Gaussian stationary, such that
 $\langle \xi(t) \rangle = 0$, $\langle \xi(t) \xi(t') \rangle = \begin{cases} \sigma^2, & t = t' \\ 0, & t \neq t' \end{cases}$